

Föreläsning 21/11-13

Lemma 6.3.5

If the state space S is finite (= finite number of possible values for Markov chain). Then at least one state is persistent and all persistent states are non-null.

Proof: $1 = \sum_j P_{ij}(n)$

If j is transient then $\sum_{n=1}^{\infty} P_{ij}(n) < \infty \Rightarrow P_{ij}(n) \rightarrow 0$ as $n \rightarrow \infty$

If all j is transient then $P_{ij}(n) \rightarrow 0$ as $n \rightarrow \infty$ for every j .

$\lim_{n \rightarrow \infty} 1 = \lim_{n \rightarrow \infty} \sum_j P_{ij}(n) = \sum_j \lim_{n \rightarrow \infty} P_{ij}(n) = 0$ contradiction

\Rightarrow at least one j is persistent. \blacksquare

$S = T \cup C_1 \cup C_2 \dots$

T transient states

C_1 closed set of persistent states that intercommunicate

C_2

\vdots

Thm 6.3.2 in each C_n all states are either null or non-null.

Assume that one of the C 's, say C_n , have null states,

$\mu_j = \infty$ all $j \in C_n$

Thm 6.2.9 says that $P_{ij}(k) \rightarrow 0$ as $k \rightarrow \infty \forall j \in C$.

$$1 = \sum_{\substack{\text{all } j \\ (i \text{ in } C_n)}} P_{ij}(k) = \sum_{j \in C_n} P_{ij}(k)$$

$$\lim_{k \rightarrow \infty} 1 = \lim_{k \rightarrow \infty} \sum_{j \in C_n} P_{ij}(k) = \sum_{j \in C_n} \lim_{k \rightarrow \infty} P_{ij}(k) = 0 \text{ contradiction as before. } \blacksquare$$

6.5 Reversibility

X_n Markov chain, started according to it's stationary dist.

(which is assumed to exist).

Consider time reversed process $Y_n = X_{N-n}$ for some (big) N .

$$\begin{array}{c} X_n \xrightarrow{n} N \\ Y_n \xleftarrow{n} N \end{array}$$

• Is Y_n also a Markov chain?

Check: $P(Y_{n+1} = i_{n+1} | Y_n = i_n, \dots, Y_0 = i_0) = \frac{P(X_{N-n+1} = i_{n+1}, Y_n = i_n, \dots, Y_0 = i_0)}{P(Y_n = i_n, \dots, Y_0 = i_0)} =$

$$= \frac{P(X_N = i_0, X_{N-1} = i_1, \dots, X_{N-n+1} = i_{n+1})}{P(X_N = i_0, X_{N-1} = i_1, \dots, X_{N-n} = i_n)} = \frac{M_{i_{n+1}}^{(N-n+1)} P_{i_{n+1} i_n} \dots P_{i_2 i_1} P_{i_1 i_0}}{M_{i_n}^{(N-n)} P_{i_n i_{n-1}} \dots P_{i_2 i_1} P_{i_1 i_0}} =$$

$$= \frac{\pi_{i_{n+1}} P_{i_{n+1} i_n}}{\pi_{i_n}} \Rightarrow P(Y_{n+1} = j | Y_n = i, \dots) = \frac{\pi_j P_{ji}}{\pi_i} \Rightarrow \text{Markov!}$$

• When does \mathbb{Y}_n have same transition matrix as \mathbb{X}_n ?

def. when that is so we call \mathbb{X}_n reversible!

Check: Happens when $P_{ij} = \frac{\pi_j P_{ji}}{\pi_i} \iff \pi_i P_{ij} = \pi_j P_{ji}$ all i and j .

Thm

Suppose we have a row matrix π which is a PMF such that $\pi_i P_{ij} = \pi_j P_{ji}$ all i, j . The π is stationary distribution for \mathbb{X} and \mathbb{Y} is reversible.

Proof: Prove that $\pi P = \pi$, $(\pi P)_k = \sum_j \pi_j P_{jk} = \sum_j \pi_j P_{kj} = \pi_k$ \square

Is all about continuous Markov chains.

$P(\mathbb{X}_{t_{n+1}} = x_{n+1} | \mathbb{X}_{t_n} = x_n, \dots, \mathbb{X}_{t_0} = x_0) = P(\mathbb{X}_{t_{n+1}} = x_{n+1} | \mathbb{X}_{t_n} = x_n)$
for discrete valued continuous time process $(\mathbb{X}_t; t \geq 0)$.